

Time to Event Prediction Based on Risk Calculation from Longitudinal Biomarkers using Bayesian Hierarchical Changepoint Mixture Models

Lynette M. Smith¹, Morshed Alam¹, Sukhwinder Kaur² ¹Department of Biostatistics, ²Department of Biochemistry and Molecular Biology University of Nebraska Medical Center, Omaha, NE, USA

Purpose/Aims

- Develop a Bayesian changepoint mixture model on longitudinally measured biomarkers for prediction of overall survival in early pancreatic cancer.
- □ The newly developed method will utilize serial marker measurements in combination with patient characteristics to improve biomarker performance in predicting pre- or post-treatment survival outcome.

Background

Pancreatic cancer (PC) is an extremely aggressive malignancy with a 5-year overall survival (OS) rate of <8%. Efforts have been made to develop methodologies for diagnostic tests, but there is also a need for biostatistical methods that utilize serially measured biomarkers to enable improved prediction of OS pre- and post-treatment following resection. Longitudinally measured biomarkers can show a variety of trends over time. The rate of change in biomarker levels in combination with the absolute expression level could better predict presence/progression of disease and OS. Currently, few methods incorporate a joint model for OS that could capture these changes in cancer patients. The growing interest in personalized medicine as well as tailored decision making on patient conditions has popularized dynamic prediction in statistical joint modeling with longitudinal and survival outcomes. Longitudinal biomarkers could be utilized in pancreatic cancer to predict overall survival for individuals to aid in decision making. We hypothesize that the use of longitudinal biomarkers in a changepoint mixture model will improve survival prediction in PC.

Analysis Methods

Mixture Model:

We will model the longitudinal biomarker data and OS in a Bayesian framework. $y_{ii}(t)$ is the biomarker data for patient *i* at time *j*.

Change point model: For cases that produce the biomarker, we will model the biomarkers as a piecewise linear function, with

$$E(Y_{ij}|t_{ij}, I_i = 1) = \theta_i + \gamma_i(t_{ij} - \tau_i),$$

where t_{ij} is the age of patient *i* when they enter the study at time $j = 0, ..., k_i$. I_i is an unobserved indicator included to distinguish cases that produce extra levels of the biomarker to those that do not. θ_i is the mean level of the biomarker for patient i and γ_i is the slope of the biomarker, which begins at the unobserved change point τ_i .

Constant model: We assume the biomarker is relatively constant for cases that do not produce extra levels of the biomarker, shown in this model.

$$E(Y_{ij}|t_{ij}, I_i = 0) = \theta_i$$

We will investigate multiple priors for these parameters before making the final decision. Normal priors will be used for the biomarker distributions, as a starting point.

For event time data, the Cox proportional hazards model is assumed:

$$\lambda_0(t_i) \exp[\beta' x(t_i, \theta_i, \gamma_i, \tau_i)]$$

 $\lambda_0(t_i)$ is the baseline hazard function, β is a vector of regression coefficients and x contains the biomarker information. We will apply Bayesian model fitting and testing using Markov chain Monte Carlo (MCMC). The method will be tested in the PC dataset as well as in a simulated dataset. Goodness of fit will be assessed by Q-Q plots as in Skates. We will also examine the sensitivity to prior distribution selection by choosing multiple priors and comparing the posterior summaries for the chosen priors.

- Determine patient biomarker profiles (see figure 1). I. Flat or
 - Increasing or decreasing slope with a changepoint.
- Combine patients into groups based on biomarker profile These groups of profiles are the "mixture" part of the model.
- Fit a joint model of survival and the longitudinal biomarker profiles.

- 43 early stage pancreatic cancer patients treated at UNMC from 2002-2011¹.
 - Demographic data: age, sex, race, alcohol smoking history.
 - Overall survival information.
 - Longitudinally measured biomarkers, MUC5AC and CA19-9.

- Table 1 shows the patient characteristics. Most are stage II, male, show moderate differentiation, and have a family history of PC.
- Patient profiles are shown in Figure 2 for MUC5AC and CA scale. CA19-9 shows mostly flat profiles whereas MUC5AC changes in biomarker levels. They produce a challenge in the when biomarkers are collected are not uniform.
- Overall survival distribution of the patients. The median OS is 1.8 years in these early stage PC patients (figure 2).
- Figure 3 shows the correlation between the biomarkers and histograms of their distributions. MUC5AC and CA19-9 show a positive correlation.

CA19-9 shows generally flatter profiles, c) overall survival distribution





Figure 3. Log biomarkers (MUC5AC and CA19-9) are positively correlated (r=0.52). Histogram for both log MUC5AC and log CA19-9 depict that the distributions are not normally distributed.







A19-9 on the log
C shows extreme
at time points

Table 1. Patient characteristics (n=43)				
		N(%)		
Age, years	Median (range)	67 (46-83)		
Gender	Male Female	28 (65%) 15 (35%)		
Stage	I II	6 (14%) 37 (86%)		
Grade	Moderate Poor Well Unknown	28 (65%) 9 (21%) 3 (7%) 3 (7%)		
Smoking status	Current Former Never	15 (35%) 14 (33%) 14 (33%)		
Alcohol use	Yes	24 (57%)		
Family history of PC	Yes	33 (79%)		
Diabetes	Yes	15 (35%)		









Translational Research

Results cont.

Additionally, we fit a frailty model of OS looking at the longitudinally measured biomarkers, accounting for correlation within person. Here we used post surgical measurements only.

Table 2. Overall survival with two biomarker predictors

			A
	HR	95% CI	P-value
n(CA19-9)	1.27	1.06-1.51	0.0098
n(MUC5AC)	0.95	0.90-1.0	0.061

Higher levels of CA19-9 following surgery are significantly associated increased risk of death. Lower levels of MUC5AC are

marginally associated with increased risk of death.

Next steps

• Classify biomarker profiles into patterns.

• Fit the mixture model, determining how to implement the changepoints.

• Determine the fit of the model.

• Compare to a simpler model without the mixture component.

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Contact information

Lynette M. Smith, PhD

Assistant Professor, Department of Biostatistics

College of Public Health, University of Nebraska Medical Center

Phone: (402) 559-8114

email:

lmsmith@unmc.edu

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